# MS&E 125: Intro to Applied Statistics Inference and confidence intervals

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April 28, 2023

# Announcements

- Quizzes will be every other week (weeks 2,4,6,8,10) either in-class on W or take-home on F
- Section today will involve candy...
- Info posted to website on projects

# **Confusions from last time**

- z-score
- $\blacktriangleright \theta$  vs  $\hat{\theta}$
- estimator
- standard error
- confidence interval
- CLT
- analytic vs computational solution

### **Pearson vs Fisher**

- what is a distribution?
- parameters
- statistics aka estimators

source: David Salzburg, "The Lady Tasting Tea: How Statistics Revolutionized Science in the Twentieth Century"

### Models and samples

# a **statistical model** says how data is generated example: we model a coin flip as a Bernoulli random variable with parameter $\theta$ we can **sample** from that model to create a dataset example: $X_1, \ldots, X_n \sim \text{Bernoulli}(\theta)$

### **Application: process control**

- Intel produces microprocessors and integrated circuits with varying performance using a complex process.
- Chips are binned based on performance, with higher-performing chips assigned to higher grades.
- Lower-performing chips are assigned to lower grades and sold as lower-end models.
- Binning chips allows manufacturers to maximize their process and offer customers a range of performance options.



# Outline

Models and inference

Normal approximation

Confidence intervals

# Inference

**inference** goes backwards: we use the data to make statements about the model

also called learning the model or distribution

example: we can learn the parameter  $\theta$  from the data one important kind of inference is **estimation**: we use the data to estimate some parameter of the model

- e.g., a mean or variance
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# **Q:** how to estimate $\theta$ from $X_1, \ldots, X_n$ ? **A:** $\hat{\theta} = \frac{1}{n} \sum_{i=1}^n X_i$

# Bias of an estimator

# Definition

the **bias** of estimator  $\hat{\theta}$  is  $\mathbb{E}[\hat{\theta}] - \theta$ 

the estimator is **unbiased** if  $\mathbb{E}[\hat{\theta}] = \theta$ 

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**Q:** What is the bias of  $\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} X_i$ ? **A:** 

$$\hat{\theta} - \theta = \frac{1}{n} \sum_{i=1}^{n} X_i - \theta = \frac{1}{n} \sum_{i=1}^{n} (X_i - \theta)$$
$$\mathbb{E}[\hat{\theta}] - \theta = \frac{1}{n} \sum_{i=1}^{n} (\mathbb{E}[X_i] - \theta) = 0$$

Poll: what is another example of an unbiased estimator for  $\theta$ ?

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Poll: which of these estimators is consistent?

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► for our coin flip model, 
$$\hat{se} = \sqrt{\frac{\theta(1-\theta)}{n}}$$
  
why? **Var**  $X = \theta(1-\theta)$ , so **Var** $[\hat{\theta}] = \frac{\theta(1-\theta)}{n}$ 

#### Demo

https://colab.research.google.com/github/ stanford-mse-125/demos/blob/main/inference.ipynb

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Models and inference

Normal approximation

Confidence intervals

### **Central limit theorem**

the **central limit theorem** says that the distribution of  $\hat{\theta} = \frac{1}{n} \sum_{i=1}^{n} X_i$  is approximately normal with mean  $\theta$  and variance **Var**  $X/n = \operatorname{se}(\theta)^2$ 

$$rac{\hat{ heta} - heta}{\mathsf{se}} o \mathcal{N}(0, 1)$$

- ▶ the distribution of  $\hat{\theta}$  is approximately normal with mean  $\theta$  and standard deviation se
- also true if the standard error se is replaced by the estimated standard error sê

assumptions:

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example: for our coin flip model,  $\hat{\theta} \sim \mathcal{N}(\theta, \frac{\theta(1-\theta)}{n})$ 

# Why use a normal approximation?

- normal distribution has just two parameters
- can estimate those parameters from data
- we can use those parameters to reason about tails of distribution

define the **z-score**: the number of standard deviations away from the mean

$$z = \frac{\hat{\theta} - \theta}{\mathsf{se}}$$

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### **Confidence interval**

a **confidence interval** is an interval *C* likely to contain the parameter e.g. the  $(1 - \alpha)$  confidence interval satisfies

 $\mathbb{P}[\theta \in C] \ge 1 - \alpha$ 

C is a random variable: it depends on the data X<sub>1</sub>,..., X<sub>n</sub>
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two interpretations (e.g., for 95% confidence interval *C*):

- If we repeat the experiment, we expect C to contain θ 100(1 − α)% of the time
- if we do a bunch of different experiments, we expect the 95% confidence interval to contain the true value of  $\theta$  for 95% of the experiments

### **Confidence intervals: examples**

opinion polls:

- $49\% \pm 3\%$  think U.S. should lift Cuba embargo.
- >  $38\% \pm 3\%$  think U.S. should build more nuclear power plants.
- 16% ± 4% think St. Louis Cardinals will win the World Series.

demographic surveys:

- The average height of adult males in the United States is between 5 feet 7 inches and 5 feet 10 inches
- The average salary of software engineers in San Francisco is between \$120,000 and \$140,000

# **Confidence intervals: examples**

medical research:

average weight loss of participants in a weight loss program is between 10 and 15 pounds

operations management:

- the mean response time of the website is between 2 and 3 seconds
- the mean time to check out at the grocery store is between 2 and 3 minutes

# How to construct confidence interval?

- (today) use a normal approximation with analytic formula for standard error
- (later) use a normal approximation with bootstrap estimate for standard error
- (later) use bootstrap quantiles

### Normal approximation for confidence interval

Suppose  $\hat{\theta} \approx N(\theta, se^2)$ . Then

$$C = \left[\hat{ heta} - z_{lpha/2}\hat{ ext{se}}, \hat{ heta} - z_{lpha/2}\hat{ ext{se}}
ight]$$

is an approximate  $(1 - \alpha)$  confidence interval for  $\theta$ , where  $z_{\alpha/2}$  is the  $(1 - \alpha/2)$  quantile of the standard normal distribution.

### Confidence interval for coin flip

example: for our coin flip model, we can construct a  $100(1-\alpha)\%$  confidence interval for  $\theta$  as

$$\hat{ heta} \pm extsf{z}_{lpha/2} \hat{ extsf{se}}$$

where  $z_{\alpha/2}$  is the  $(1 - \alpha/2)$  quantile of the standard normal distribution

e.g., for  $\alpha = 0.05$ , we use  $z_{0.025} = 1.96$ 

# Calibration

A (1-lpha) confidence interval is called **calibrated** if  $\mathbb{P}[\theta \in C] \approx 1- lpha$ 

if confidence interval is too large, it's useless
if confidence interval is too small, it's wrong

# Proof that normal confidence interval is calibrated

Proof:

$$\begin{aligned} \mathsf{Pr}(\theta \in \mathsf{C}_n) &= \mathsf{Pr}(\hat{\theta}_n - z_{\alpha/2}\hat{\mathsf{se}} \le \theta \le \hat{\theta}_n + z_{\alpha/2}\hat{\mathsf{se}}) \\ &= \mathsf{Pr}(-z_{\alpha/2}\hat{\mathsf{se}} \le \theta - \hat{\theta}_n \le z_{\alpha/2}\hat{\mathsf{se}}) \\ &= \mathsf{Pr}\left(-z_{\alpha/2} \le \frac{\theta - \hat{\theta}_n}{\hat{\mathsf{se}}} \le z_{\alpha/2}\right) \\ &\approx \mathsf{Pr}\left(-z_{\alpha/2} \le Z \le z_{\alpha/2}\right) \\ &= 1 - \alpha \end{aligned}$$

- se approximates the standard deviation of  $\hat{\theta}$
- the central limit theorem says that (\heta\) is approximately normal, so the standard deviation controls the tails of the distribution

 $\implies$  CI is calibrated if number of samples *n* is large enough to justify approximations